

Model defects for nuclear data evaluation

St. Gundacker
H. Leeb

Vienna University of
Technology
Atominsttitut of the Austrian
Universities, Vienna, Austria

-) Energy range of evaluated files normally limited up to 20 MeV
-) novel nuclear technologies and radioactive waste incineration methods require extension to higher energies
-) in general nuclear data evaluation is a statistical process within BAYESIAN statistics
-) unfortunately experimental data is scarce for neutron induced reactions at higher energies
-) PRIOR becomes important due to scarcity of data
-) Model defects cause great impact on prior

Bayesian statistics:

sum rule: $p(\underline{x}|M) + p(\bar{\underline{x}}|M) = 1$

product rule: $p(\underline{x}|\underline{\sigma}M) p(\underline{\sigma}|M) = p(\underline{\sigma}|\underline{x}M) p(\underline{x}|M)$

Bayes Theorem (1763):

$$p(\underline{x}|\underline{\sigma}M) = \frac{1}{p(\underline{\sigma}|M)} p(\underline{\sigma}|\underline{x}M) p(\underline{x}|M)$$

**updated probability
distribution of the
set of parameters**

**likelihood function,
from experiment**

PRIOR

$$p(\underline{x}|\underline{\sigma}M)$$

.... probability distribution of parameters \underline{x} for a given model M and data $\underline{\sigma}$

$$p(\underline{\sigma}|\underline{x}M)$$

.... probability distribution of data $\underline{\sigma}$ for a model M with parameters \underline{x}

$$p(\underline{x}|M)$$

.... probability for the occurrence of \underline{x} when M is true

The contributions to the covariance matrix of the model are:

$$M^{(mod)} = M^{(par)} + M^{(num)} + M^{(def)}$$

**Parameter
uncertainties**

**Numerical
implementation
errors**

**Deficiencies of the model,
is of non-statistical nature**

The contributions to the covariance matrix of the model are:

$$M^{(mod)} = M^{(par)} + M^{(num)} + M^{(def)}$$

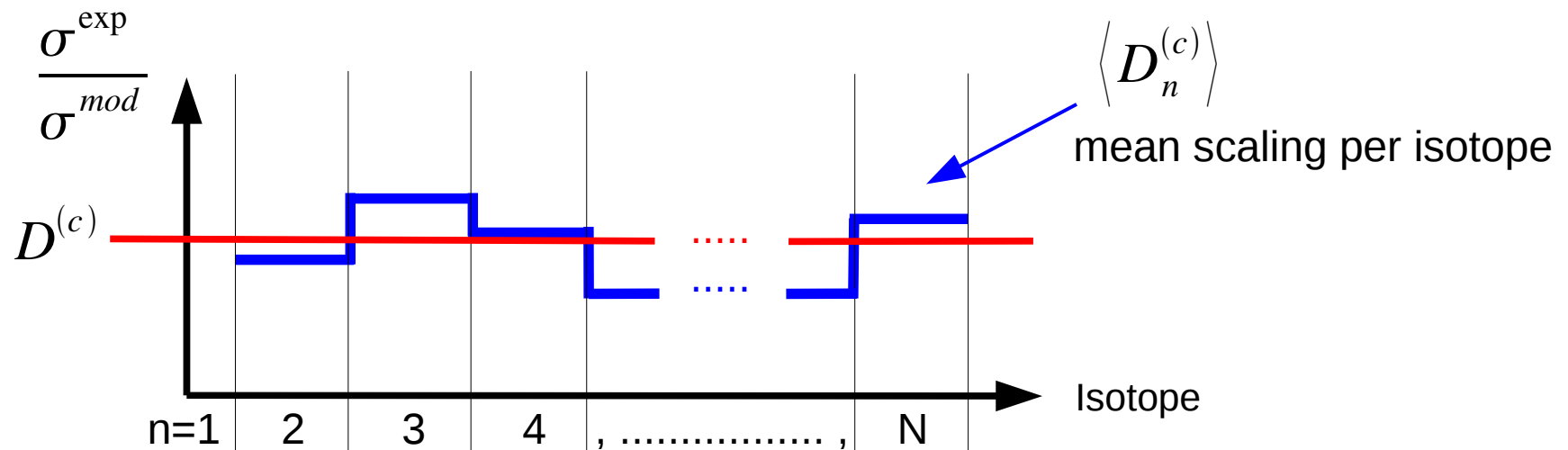
-) cannot be determined within the considered nuclear model
-) It is required to involve experimental data in the procedure.
-) Only corresponding data from neighbouring nuclei is considered.

The basic idea is to introduce an overall scaling factor:

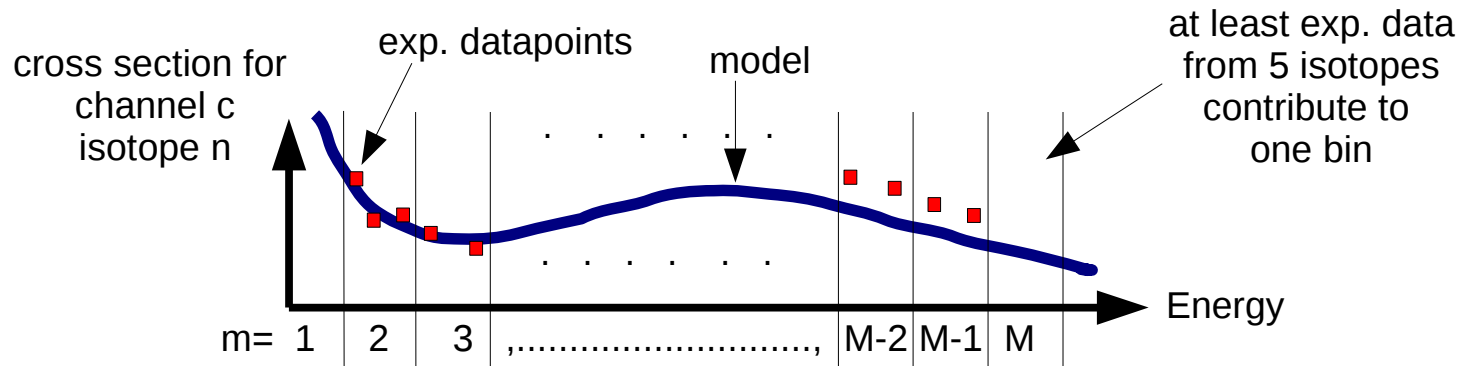
$$D^{(c)} = \frac{1}{N} \sum_{n=1}^N \langle D_n^{(c)} \rangle$$

for the reaction channel c

which is a mean quotient between experiment and theory $\frac{\sigma^{\text{exp}}}{\sigma^{\text{mod}}}$



energy independent scaling factors => deficiencies of model are directly reflected



scaling factor in the energy bin M

$$\langle D_n^{(c)}(E_m) \rangle = \sum_{j \in E_{bin}(m,n)} w_j^{(c,m,n)} \frac{\sigma_{ex}^{(c)}(E_j)}{\sigma_{th}^{(c)}(E_j)}$$

scaling factor per isotope:

$$\langle D_n^{(c)} \rangle = \sum_{m=1}^M w_m^{(c,n)} \langle D_n^{(c)}(E_m) \rangle$$

chosen weights emphasize values at the highest cross section

$$w_j^{(c,m,n)} = \frac{\sigma_{th}^{(c,n)}(E_j)}{\sum_{j' \in E_{bin}(m,n)} \sigma_{th}^{(c,n)}(E_{j'})}$$

$$w_m^{(c,n)} = \frac{\sigma_{th}^{(c,n)}(E_m)}{\sum_{m' \in E_{bin}} \sigma_{th}^{(c,n)}(E_{m'})}$$

covariance matrix due to model defects we define by:

$$\begin{aligned} \langle \Delta^{(c)}(E_m) \Delta^{(c')}(E_{m'}) \rangle &= \sigma_{th}^{(c)}(E_m) \sigma_{th}^{(c')}(E_{m'}) \\ &\cdot \frac{1}{\sqrt{N^{(c)}(E_m)} \sqrt{N^{(c')}(E_{m'})}} \sum_{n=1}^N \left\{ \left[\left(\langle D_n^{(c)}(E_m) \rangle - D^{(c)} \right) \left(\langle D_n^{(c')}(E_{m'}) \rangle - D^{(c')} \right) \right] \right. \\ &\left. + \delta_{cc'} \delta_{mm'} \left[\left\langle \left(D_n^{(c)}(E_m) \right)^2 \right\rangle - \left(\langle D_n^{(c)}(E_m) \rangle \right)^2 \right] \right\} \end{aligned}$$

-) it's an assumption

-) that formulation is of non-statistical nature!

covariance matrix due to model defects:

$$\begin{aligned}
 \left\langle \Delta^{(c)}(E_m) \Delta^{(c')}(E_{m'}) \right\rangle &= \sigma_{th}^{(c)}(E_m) \sigma_{th}^{(c')}(E_{m'}) \\
 &\cdot \frac{1}{\sqrt{N^{(c)}(E_m)} \sqrt{N^{(c')}(E_{m'})}} \sum_{n=1}^N \left\{ \left[\left(\left\langle D_n^{(c)}(E_m) \right\rangle - D^{(c)} \right) \left(\left\langle D_n^{(c')}(E_{m'}) \right\rangle - D^{(c')} \right) \right] \right. \\
 &\left. + \delta_{cc'} \delta_{mm'} \left[\left\langle \left(D_n^{(c)}(E_m) \right)^2 \right\rangle - \left(\left\langle D_n^{(c)}(E_m) \right\rangle \right)^2 \right] \right\}
 \end{aligned}$$

First term expresses systematical errors, represents the correlations

covariance matrix due to model defects:

$$\begin{aligned} \langle \Delta^{(c)}(E_m) \Delta^{(c')}(E_{m'}) \rangle &= \sigma_{th}^{(c)}(E_m) \sigma_{th}^{(c')}(E_{m'}) \\ &\cdot \frac{1}{\sqrt{N^{(c)}(E_m)} \sqrt{N^{(c')}(E_{m'})}} \sum_{n=1}^N \left\{ \left[\left(\langle D_n^{(c)}(E_m) \rangle - D^{(c)} \right) \left(\langle D_n^{(c')}(E_{m'}) \rangle - D^{(c')} \right) \right] \right. \\ &\left. + \delta_{cc'} \delta_{mm'} \left[\left\langle \left(D_n^{(c)}(E_m) \right)^2 \right\rangle - \left(\langle D_n^{(c)}(E_m) \rangle \right)^2 \right] \right\} \end{aligned}$$

$N^{(c)}(E_m)$ is the number of isotopes for which $\langle D_n^{(c)}(E_m) \rangle$ can be evaluated

chosen normalization accounts for nonstatistical nature of the formulation

covariance matrix due to model defects:

$$\begin{aligned} \left\langle \Delta^{(c)}(E_m) \Delta^{(c')}(E_{m'}) \right\rangle &= \sigma_{th}^{(c)}(E_m) \sigma_{th}^{(c')}(E_{m'}) \\ &\cdot \frac{1}{\sqrt{N^{(c)}(E_m)} \sqrt{N^{(c')}(E_{m'})}} \sum_{n=1}^N \left\{ \left[\left(\left\langle D_n^{(c)}(E_m) \right\rangle - D^{(c)} \right) \left(\left\langle D_n^{(c')}(E_{m'}) \right\rangle - D^{(c')} \right) \right] \right. \\ &\quad \left. + \delta_{cc'} \delta_{mm'} \left[\left\langle \left(D_n^{(c)}(E_m) \right)^2 \right\rangle - \left(\left\langle D_n^{(c)}(E_m) \right\rangle \right)^2 \right] \right\} \end{aligned}$$

$$\left\langle \left(D_n^{(c)}(E_m) \right)^2 \right\rangle = \sum_{j \in E_{bin}(m,n)} w_j^{c,m,n} \left(\frac{\sigma_{ex}^{(c)}(E_j)}{\sigma_{th}^{(c)}(E_j)} \right)^2$$

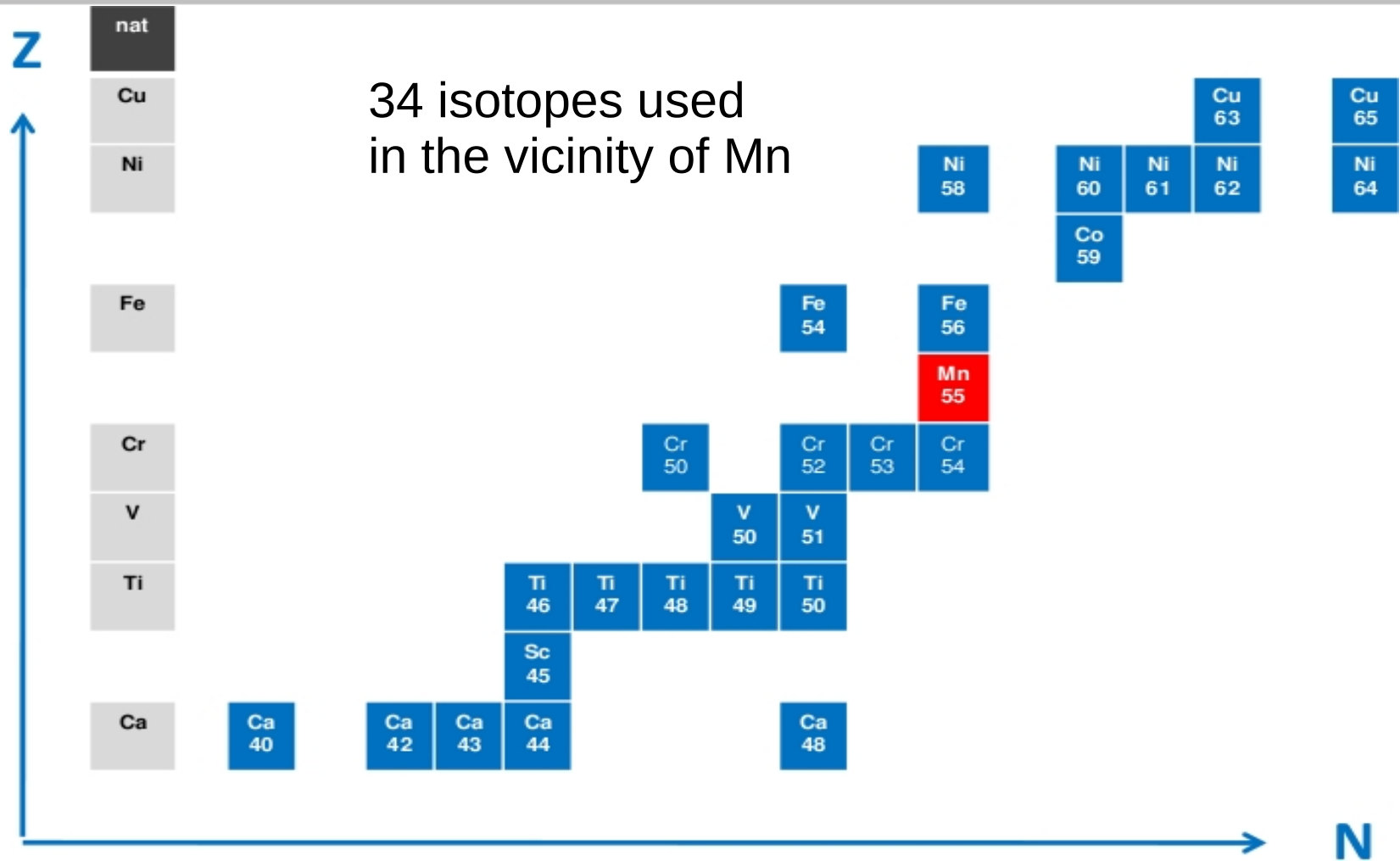
Second term is a real covariance term as defined in statistics due to fluctuations of the experimental data

covariance matrix due to model defects:

$$\begin{aligned} \langle \Delta^{(c)}(E_m) \Delta^{(c')}(E_{m'}) \rangle &= \sigma_{th}^{(c)}(E_m) \sigma_{th}^{(c')}(E_{m'}) \\ &\cdot \frac{1}{\sqrt{N^{(c)}(E_m)} \sqrt{N^{(c')}(E_{m'})}} \sum_{n=1}^N \left\{ \left[\left(\langle D_n^{(c)}(E_m) \rangle - D^{(c)} \right) \left(\langle D_n^{(c')}(E_{m'}) \rangle - D^{(c')} \right) \right] \right. \\ &\left. + \delta_{cc'} \delta_{mm'} \left[\left\langle \left(D_n^{(c)}(E_m) \right)^2 \right\rangle - \left(\langle D_n^{(c)}(E_m) \rangle \right)^2 \right] \right\} \end{aligned}$$

Correlations are defined in the usual way:

$$C^{(cc')}(E_m, E_n) = \frac{\langle \Delta \sigma^{(c)}(E_m) \Delta \sigma^{(c')}(E_n) \rangle}{\sqrt{\langle \Delta^2 \sigma^{(c)}(E_m) \rangle} \sqrt{\langle \Delta^2 \sigma^{(c')}(E_n) \rangle}}$$



channel:	n,tot	n,non	n,el	n,inl	n,2n	n,p	n, α
Nr. of isotopes	23	8	21	15	14	24	15
Nr. of data points	200162	173	2151	1666	1259	2498	772

-) optical potential (Koning and Delaroche) and level densities (CTM - TALYS) were optimised
-) because global parameters in TALYS are optimised from $A=12$ to 339
-) secures that no exp. information of ^{55}Mn goes into global parametrisation
-) we have obtained a slightly different parametrisation

lane term in neutron opt. model:

$$d_1 = 19.59 - 64.95 \frac{N-Z}{A}$$

$$d_1 = 16.0 - 16.0 \frac{N-Z}{A}$$

TALYS

level density parameters for CTM model:

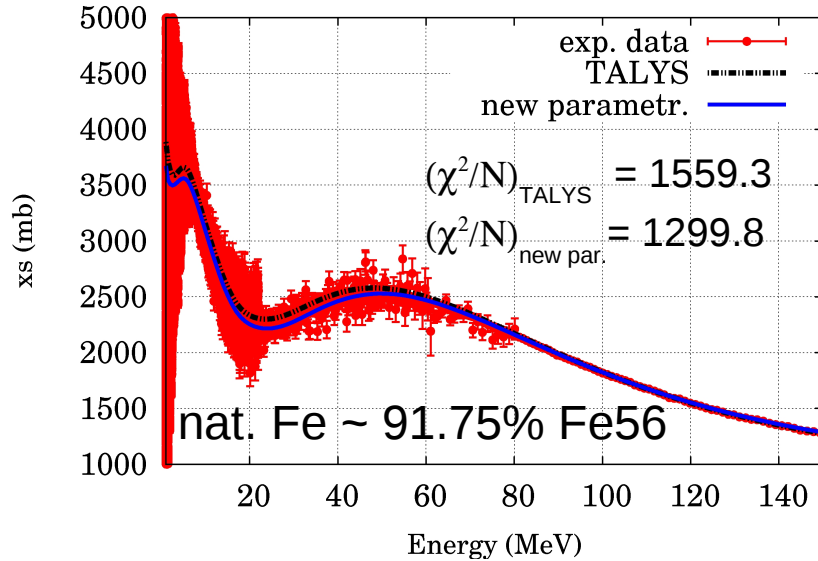
$$\alpha = 0.026220 \quad \beta = 0.270416$$

$$\gamma_1 = 0.456296$$

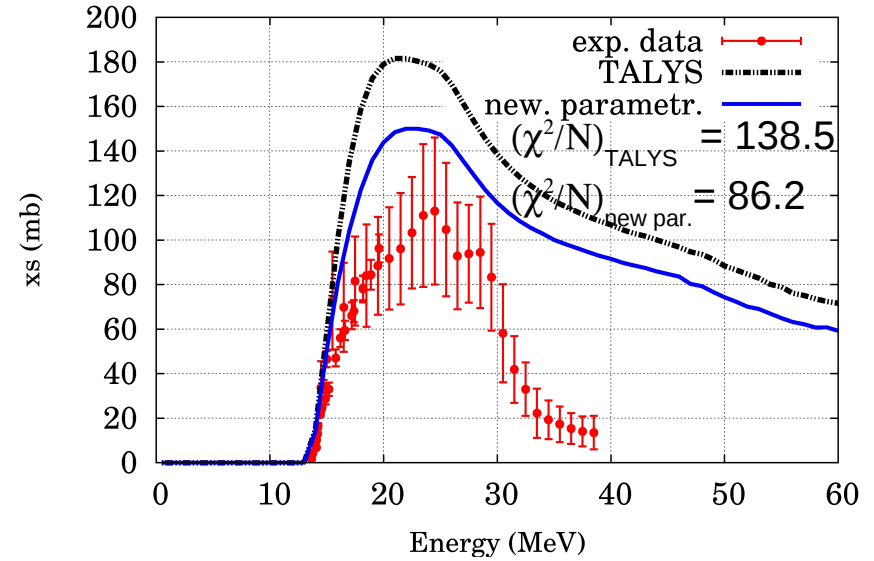
$$\alpha = 0.0207305 \quad \beta = 0.229537$$

$$\gamma_1 = 0.473625$$

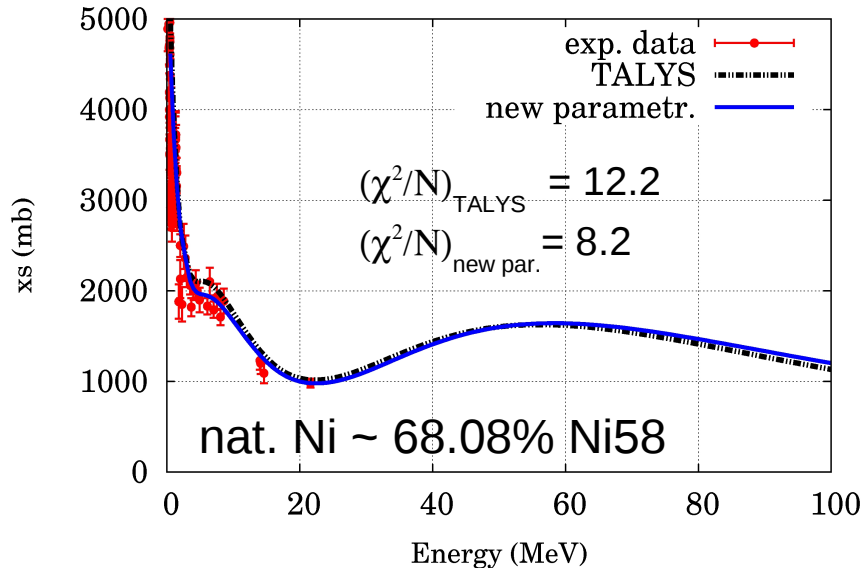
totalxs for 26Fe00



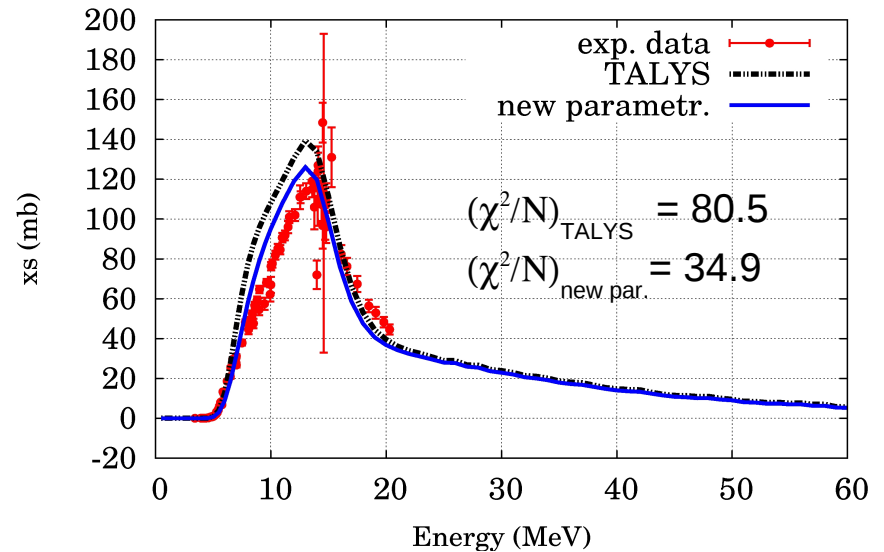
n,2n for 24Cr50



n,el for 28Ni00



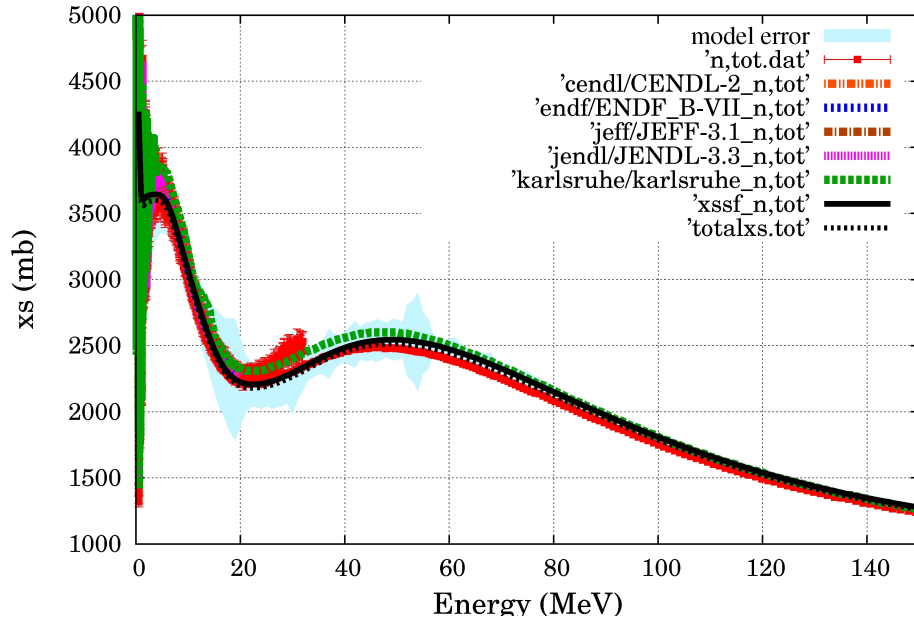
n,p for 26Fe56



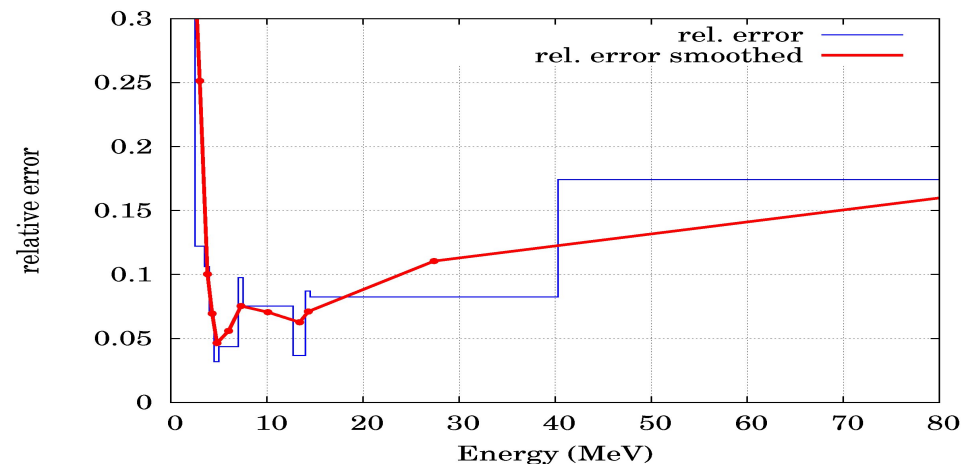
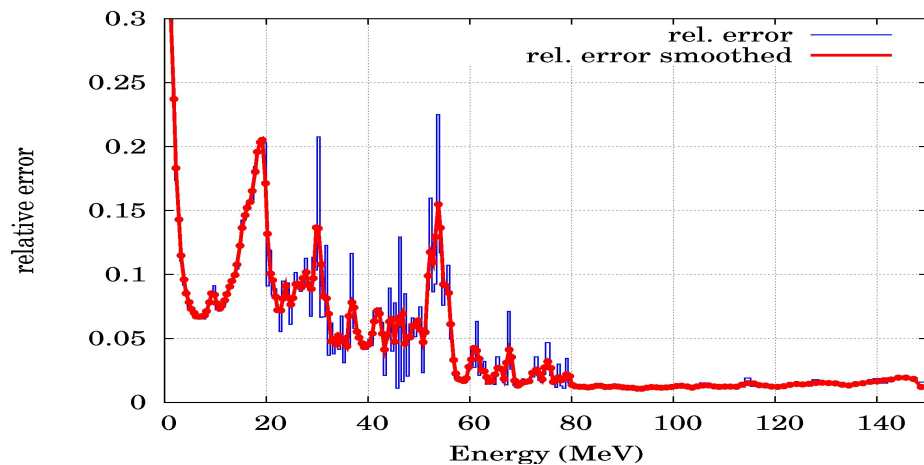
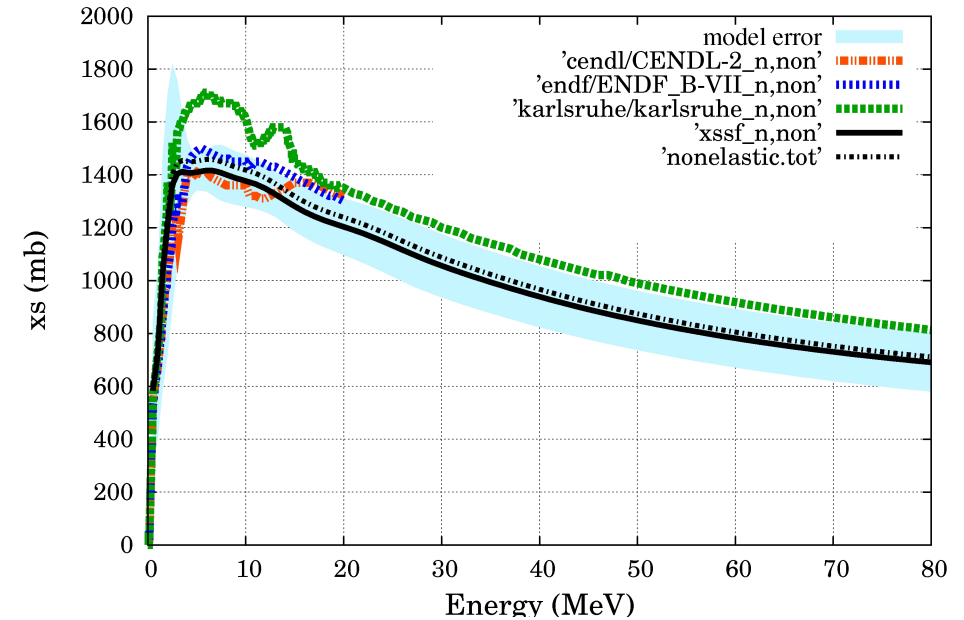
overall scaling factor (n,tot): 1.013

overall scaling factor (n,non): 0.971

comparison of n,tot

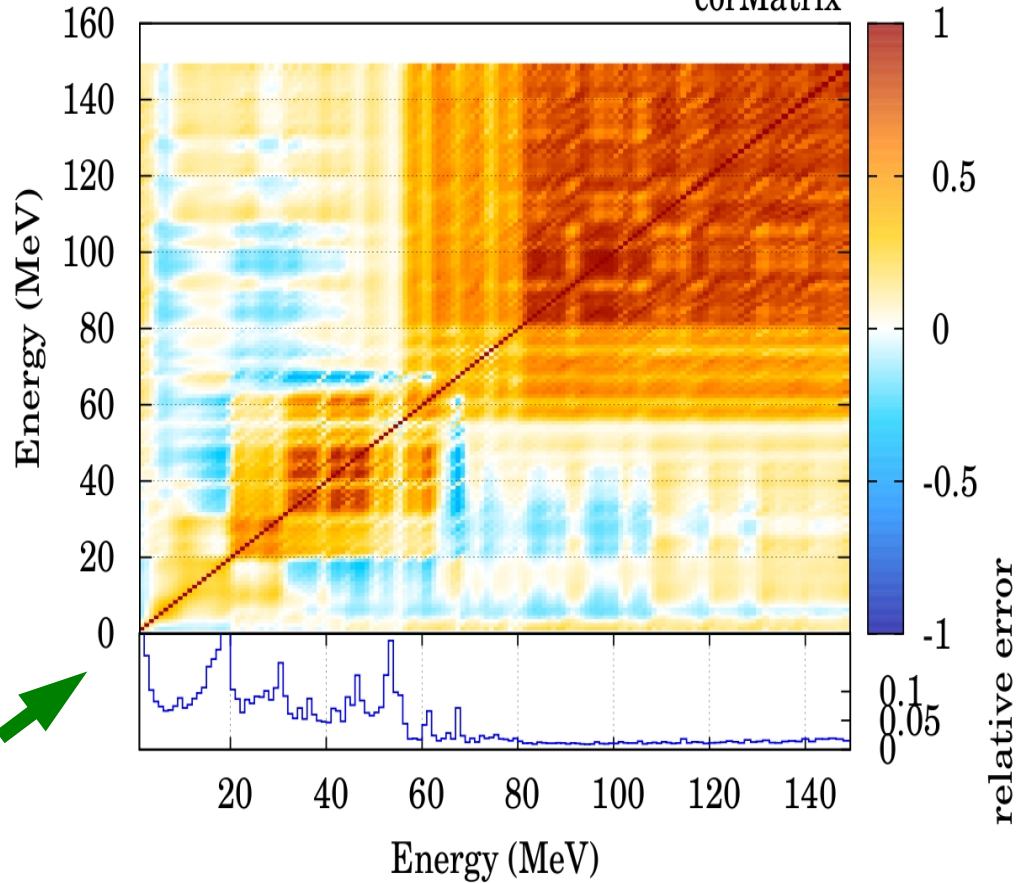


comparison of n,non



correlation (n_{tot}) with (n_{tot})

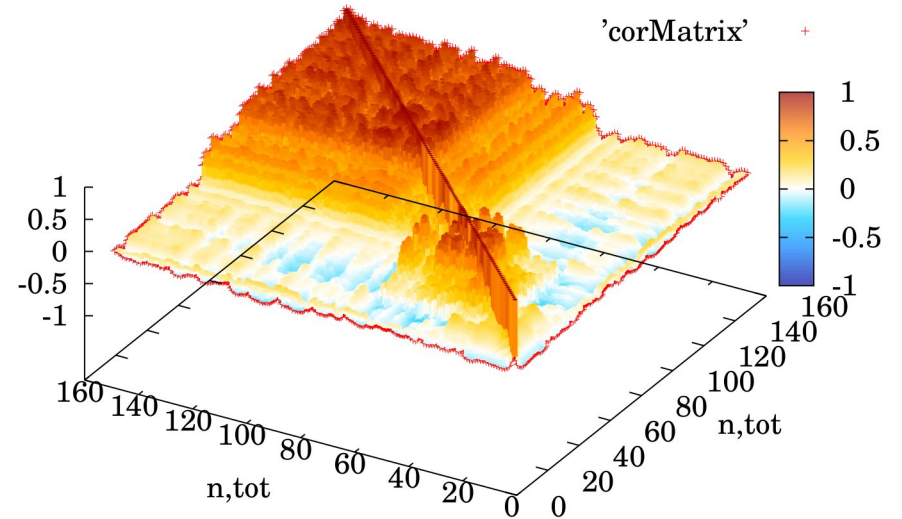
'corMatrix'



$$E_m + E_{m'} = \text{const.}$$

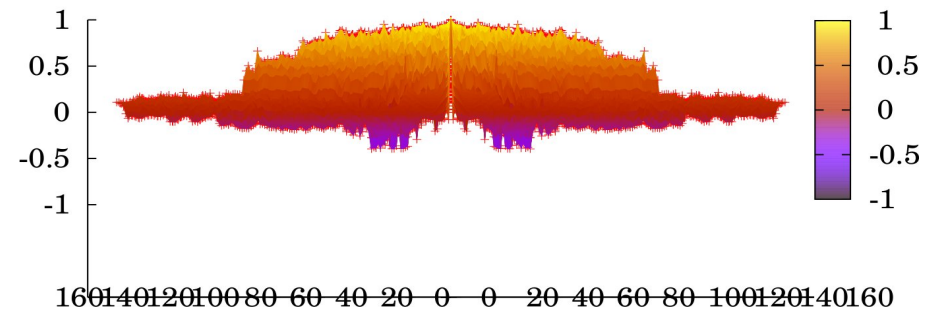
correlation matrix n_{tot} with n_{tot}

'corMatrix'



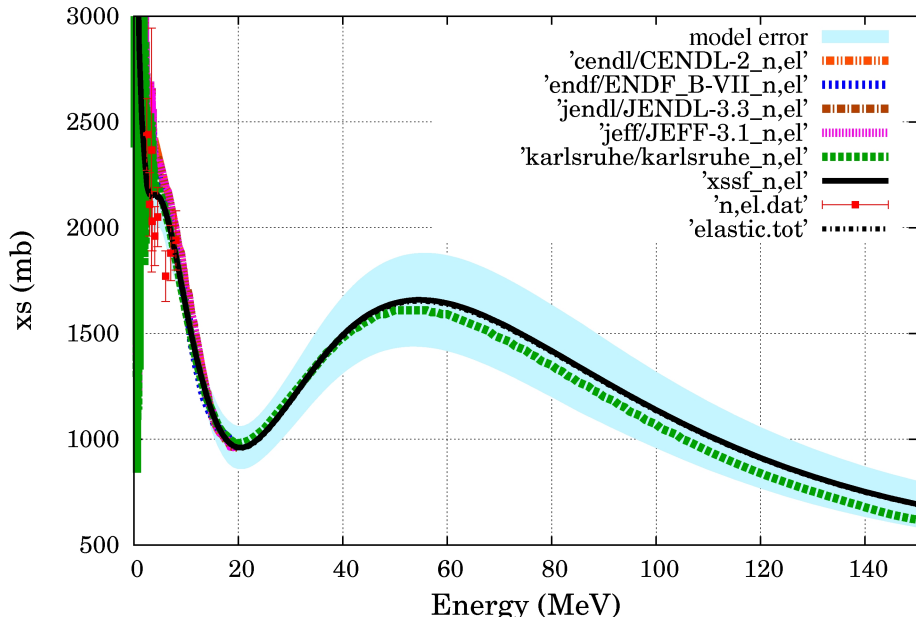
correlation matrix n_{tot} with n_{tot}

'corMatrix2'

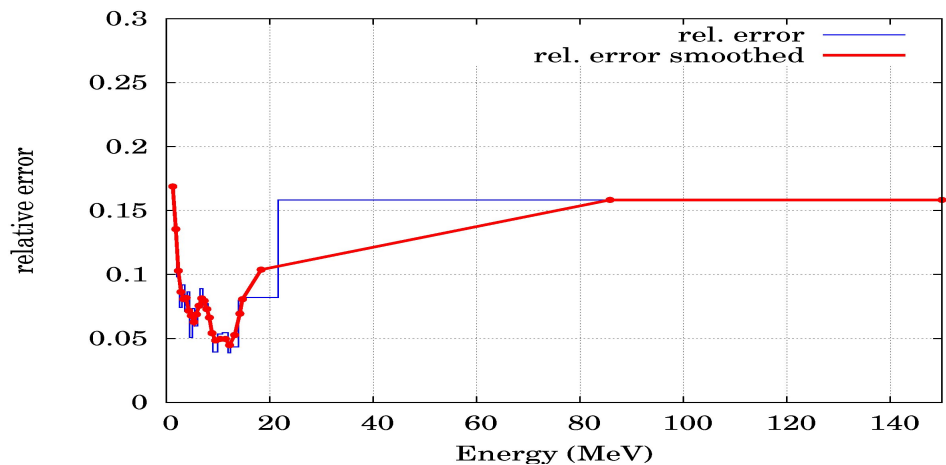
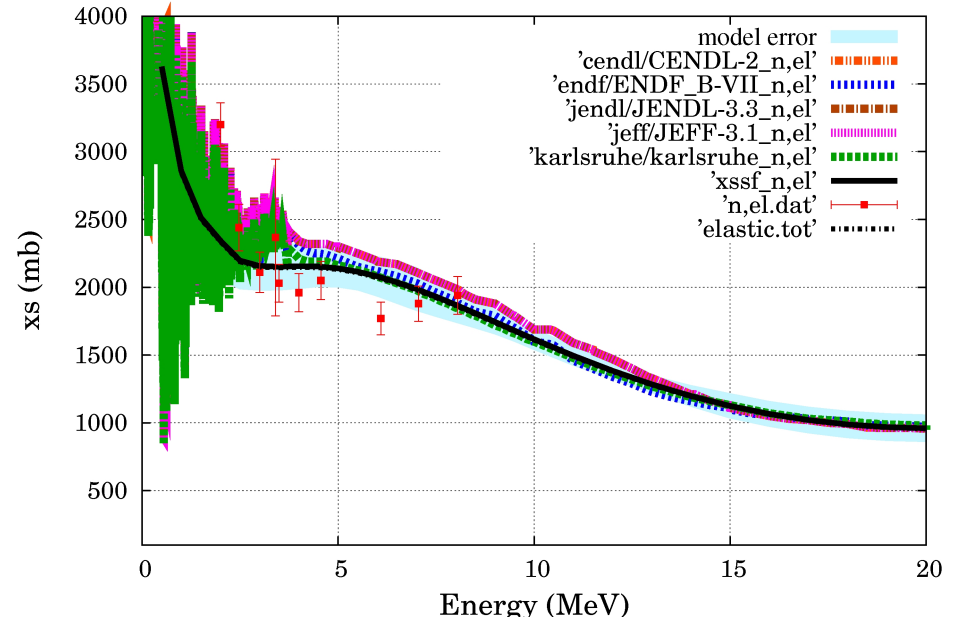


overall scaling factor (n,el): 1.004

comparison of n,el



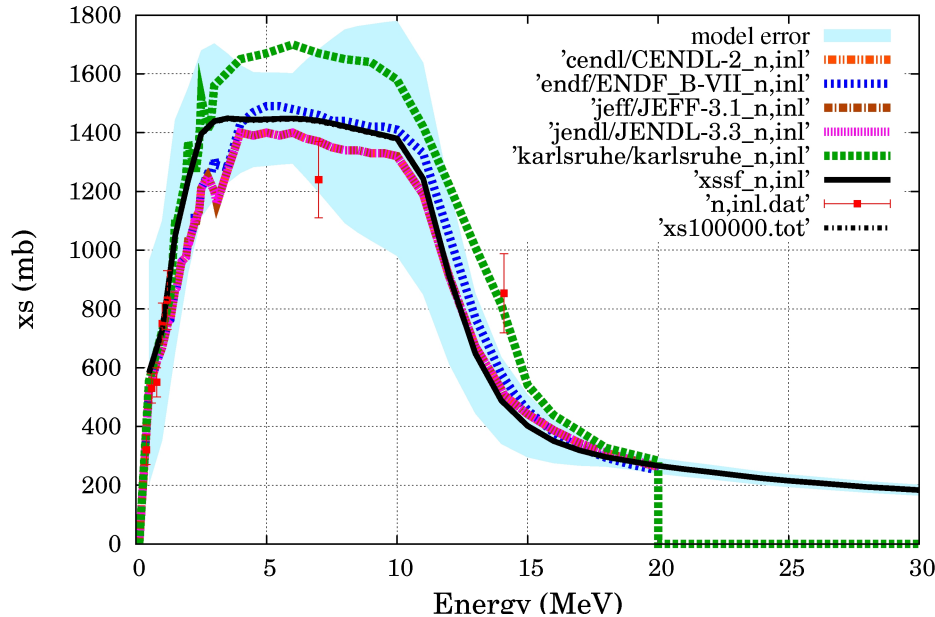
comparison of n,el



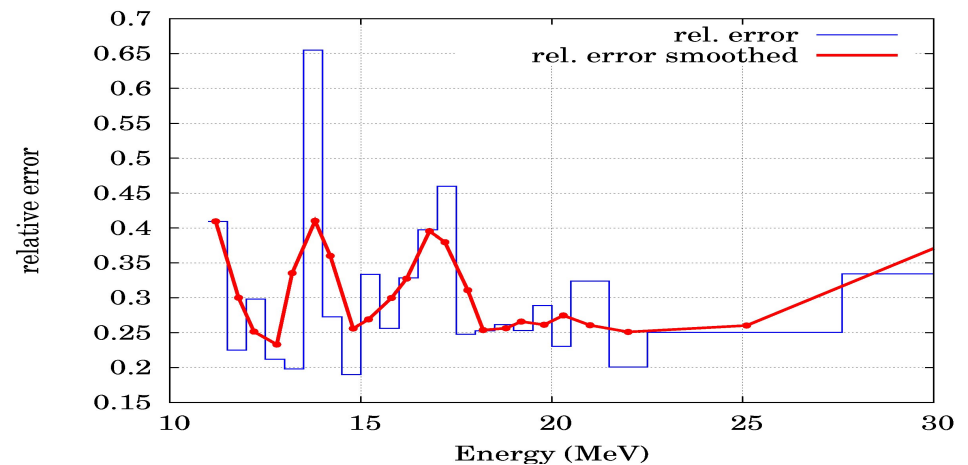
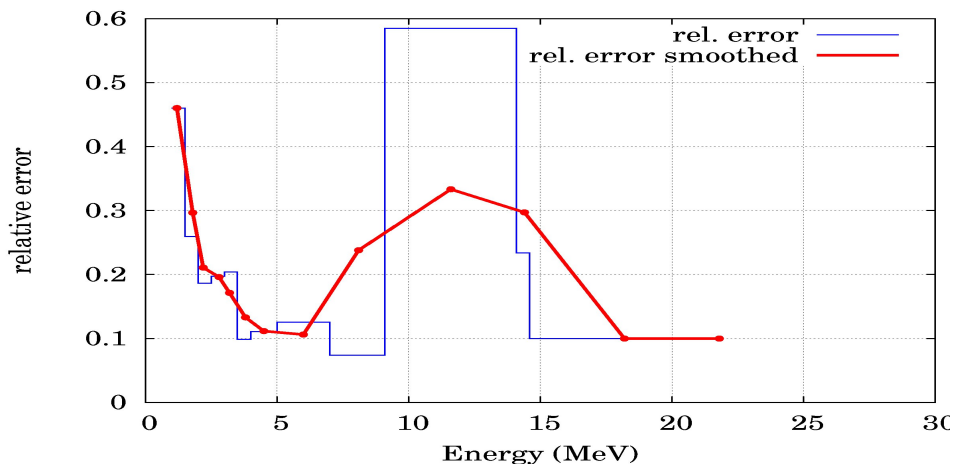
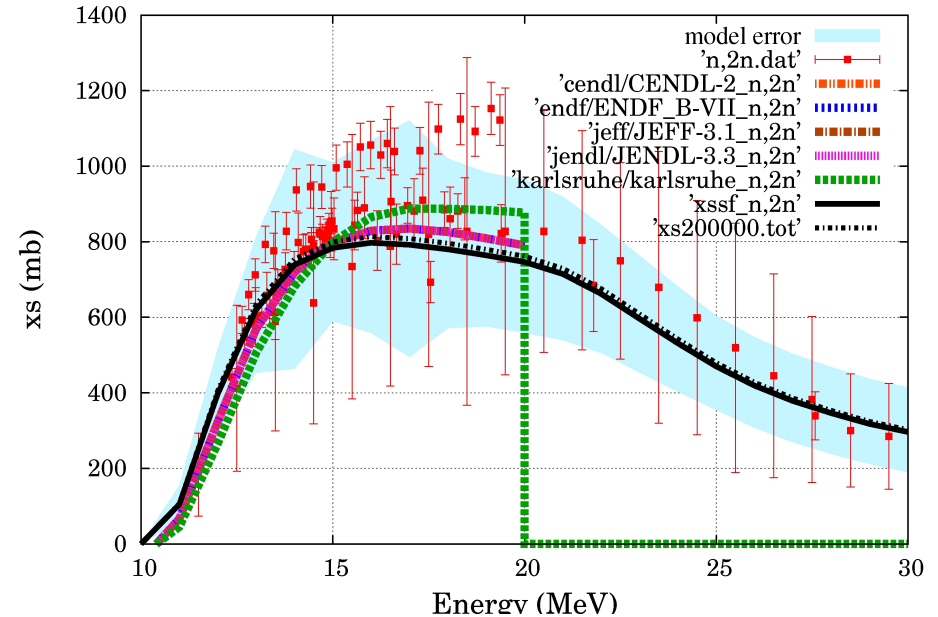
overall scaling factor (n,inl): 1.002

overall scaling factor (n,2n): 0.980

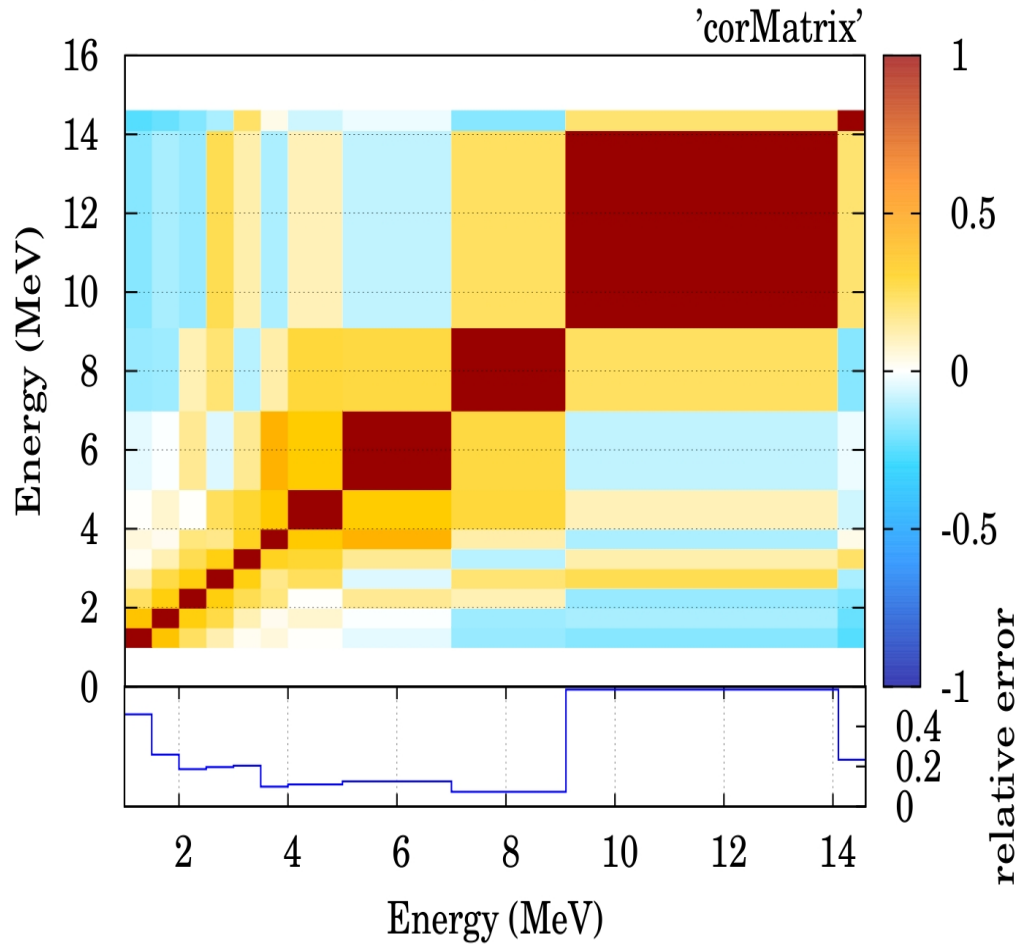
comparison of n,inl



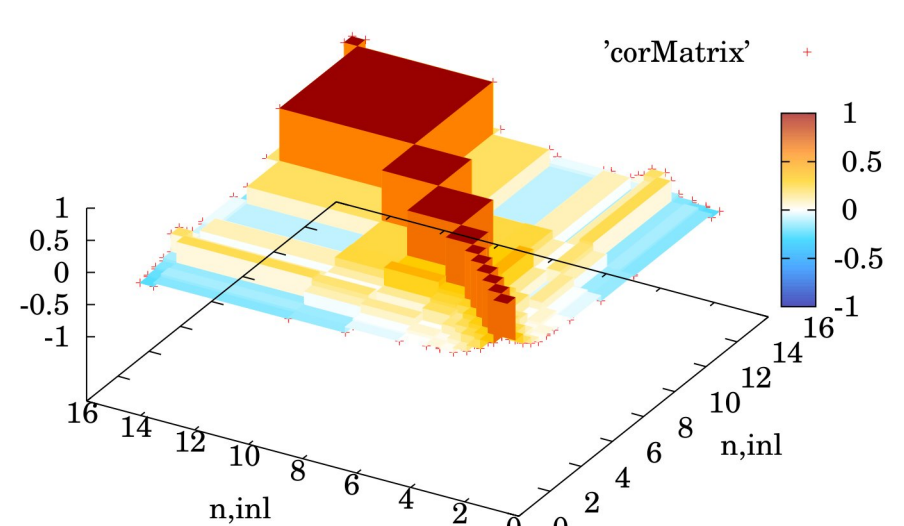
comparison of n,2n



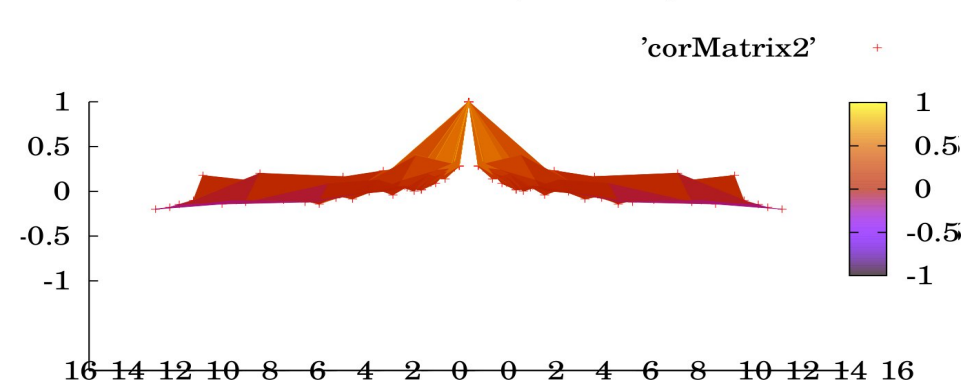
correlation ($n_{,inl}$) with ($n_{,inl}$)



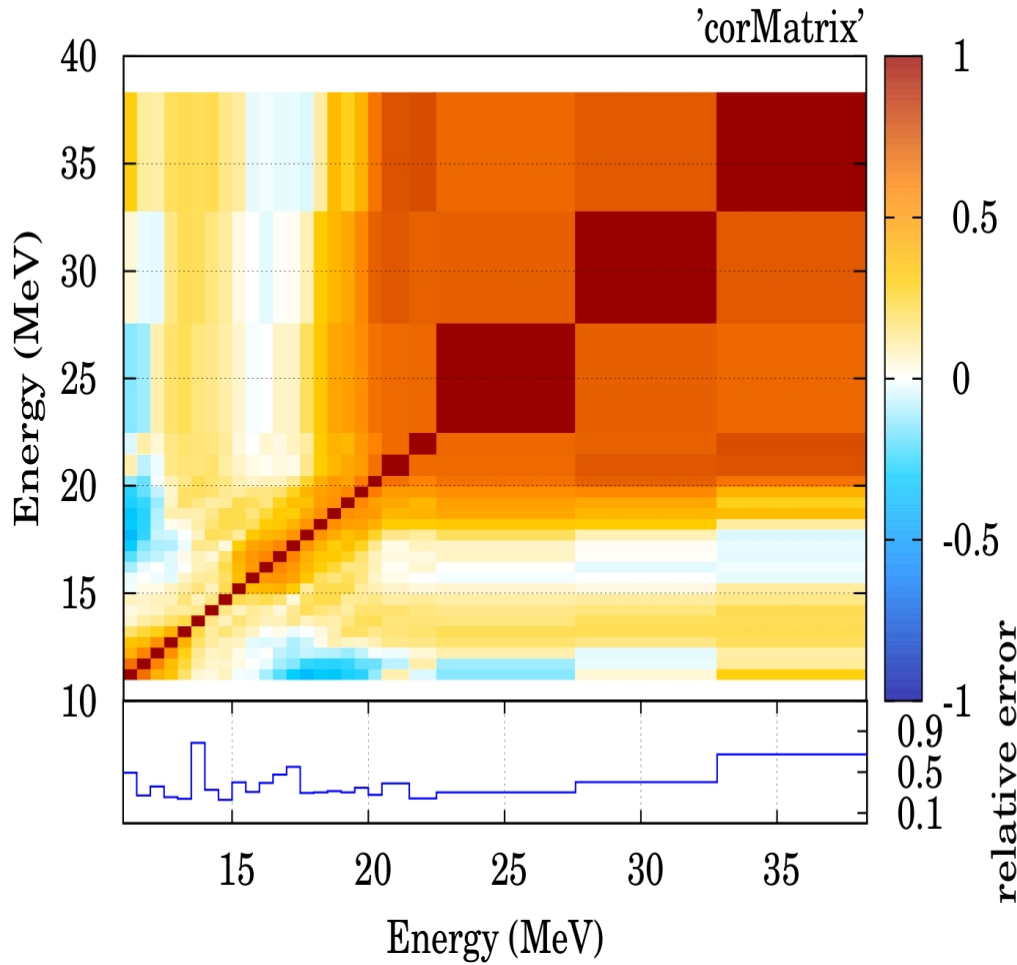
correlation matrix $n_{,inl}$ with $n_{,inl}$



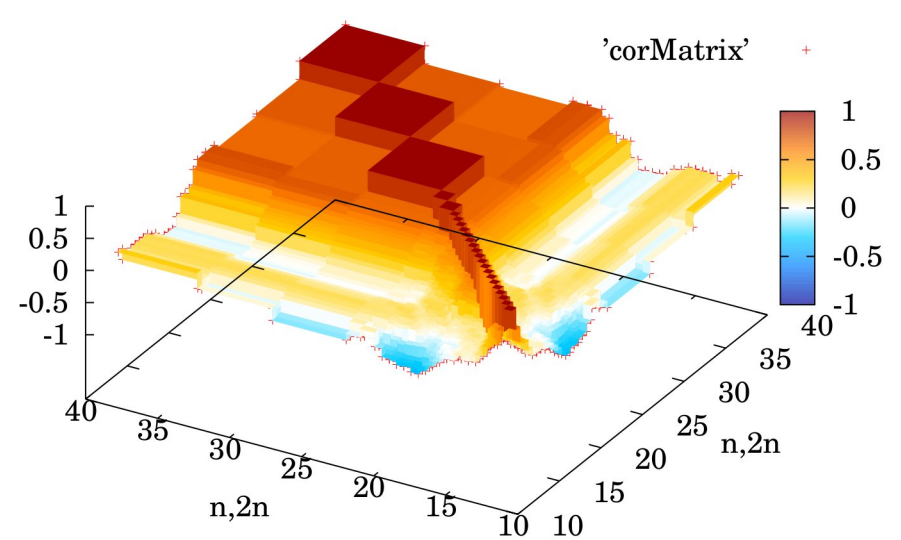
correlation matrix $n_{,inl}$ with $n_{,inl}$



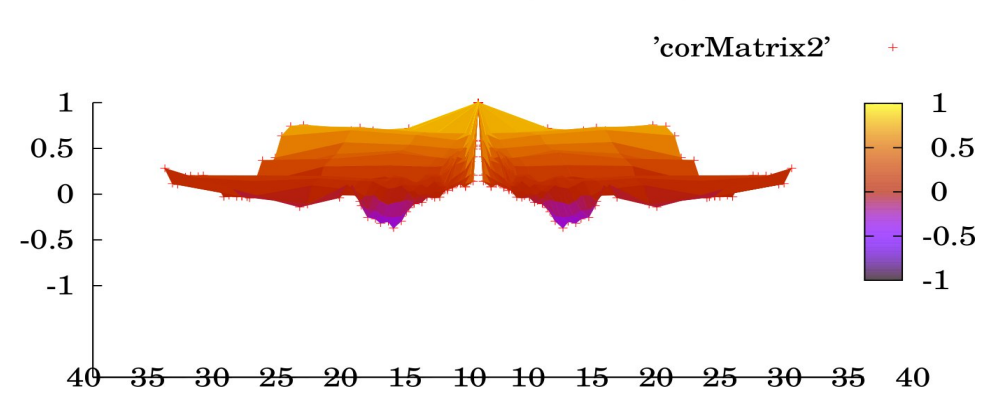
correlation (n,2n) with (n,2n)

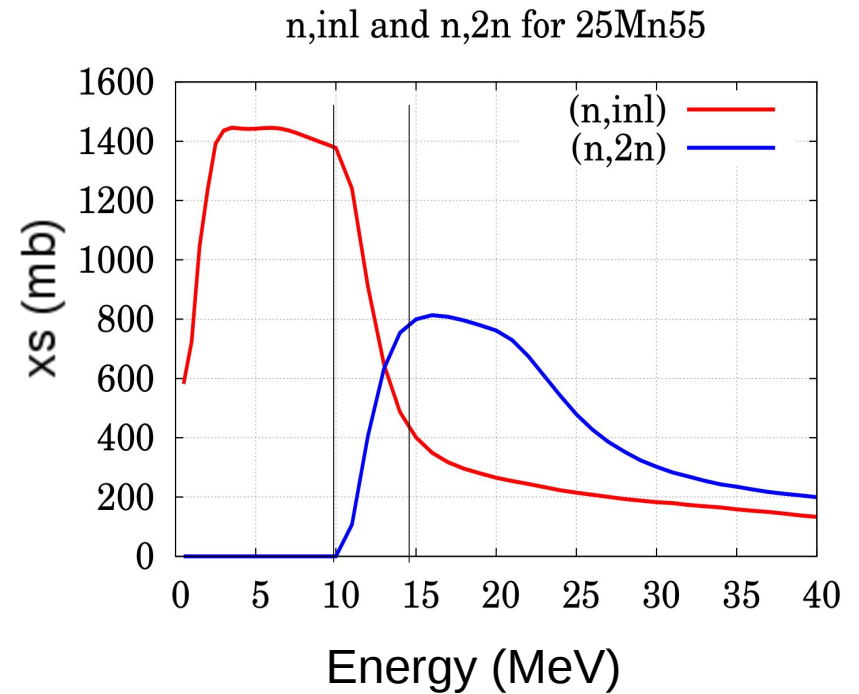
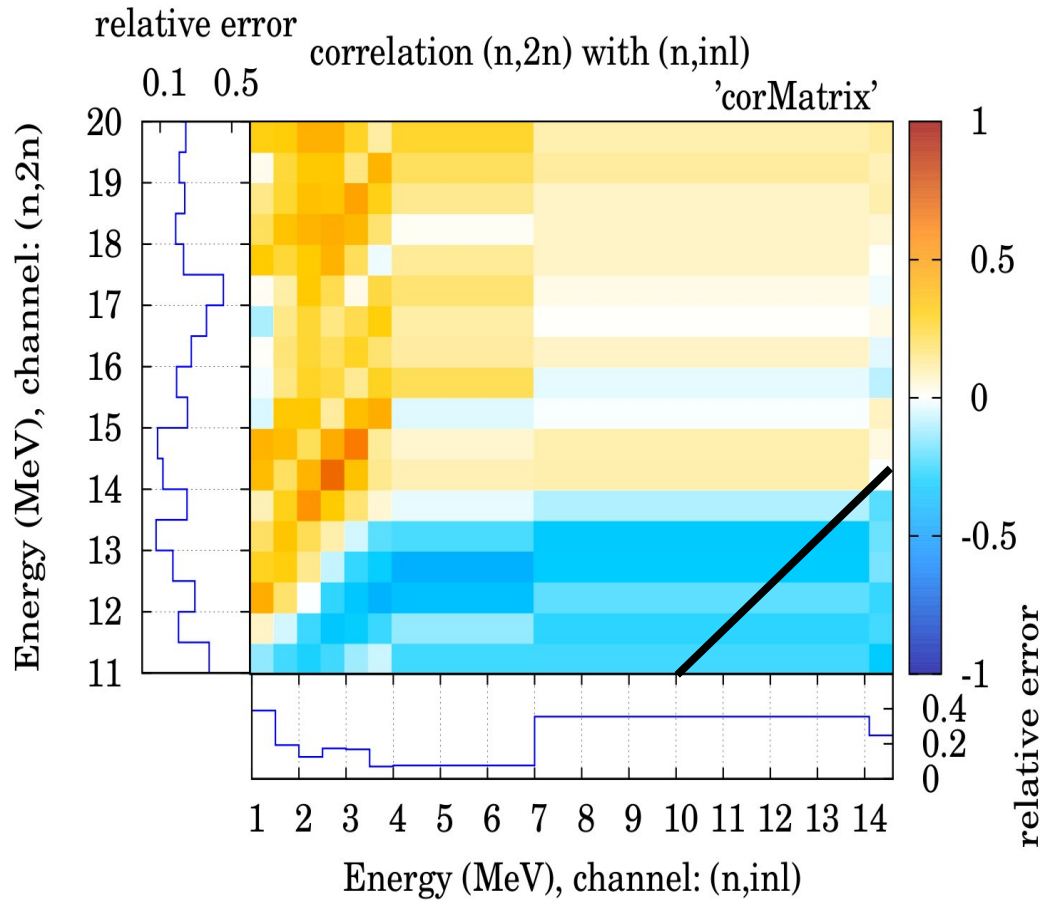


correlation matrix n,2n with n,2n



correlation matrix n,2n with n,2n

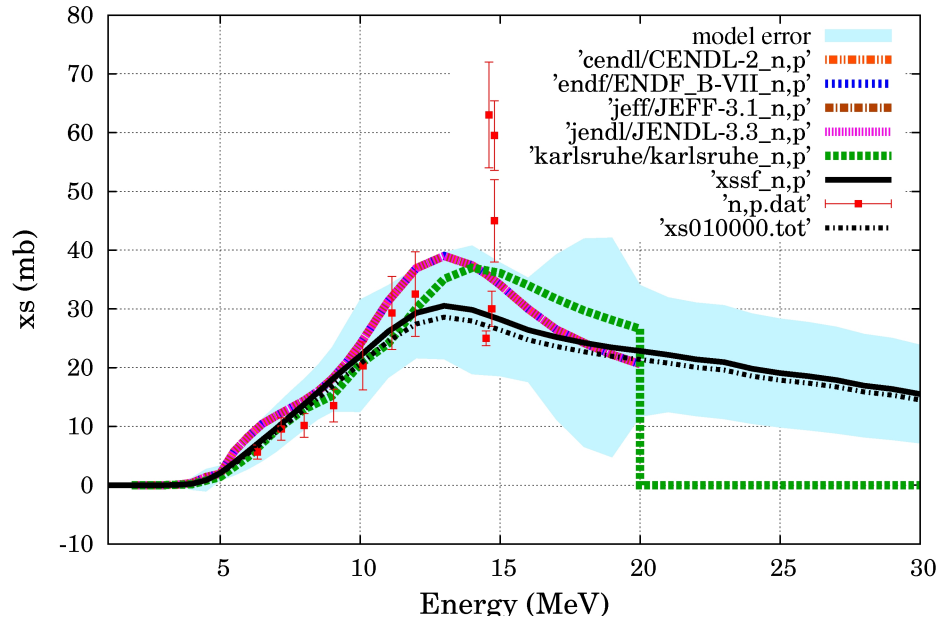




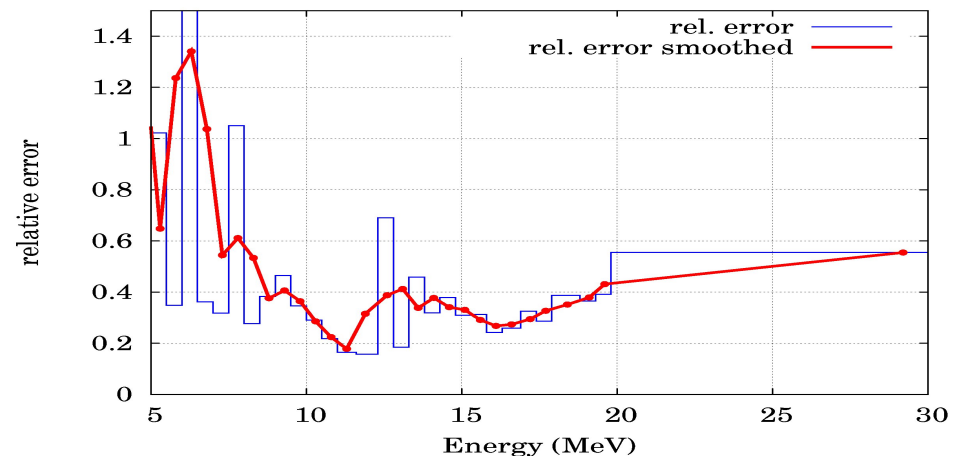
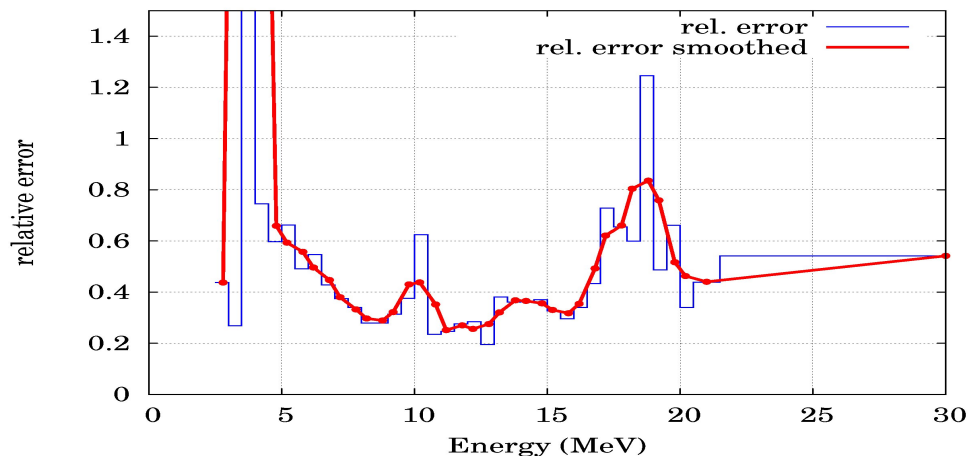
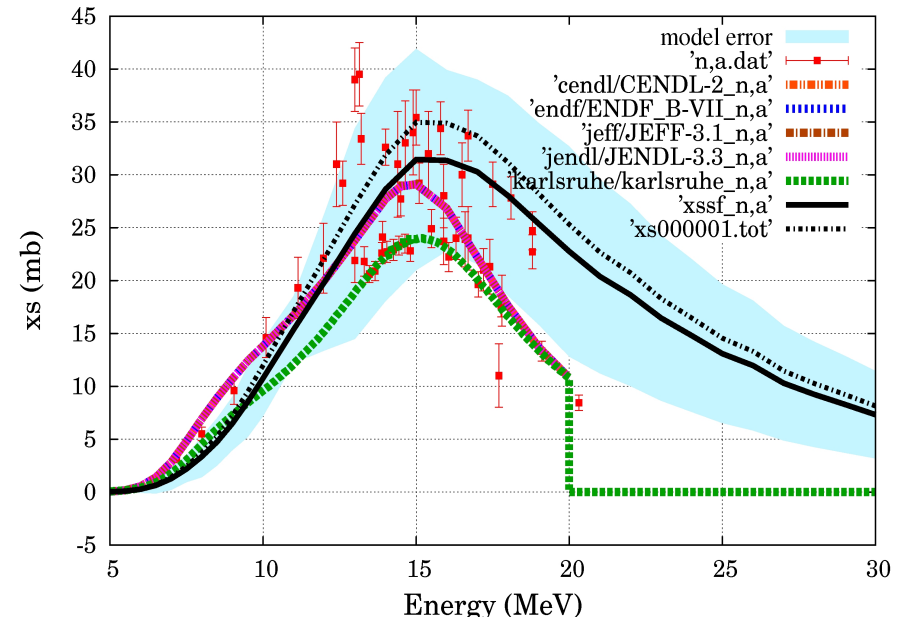
overall scaling factor (n,p): 1.068

overall scaling factor (n, α): 0.899

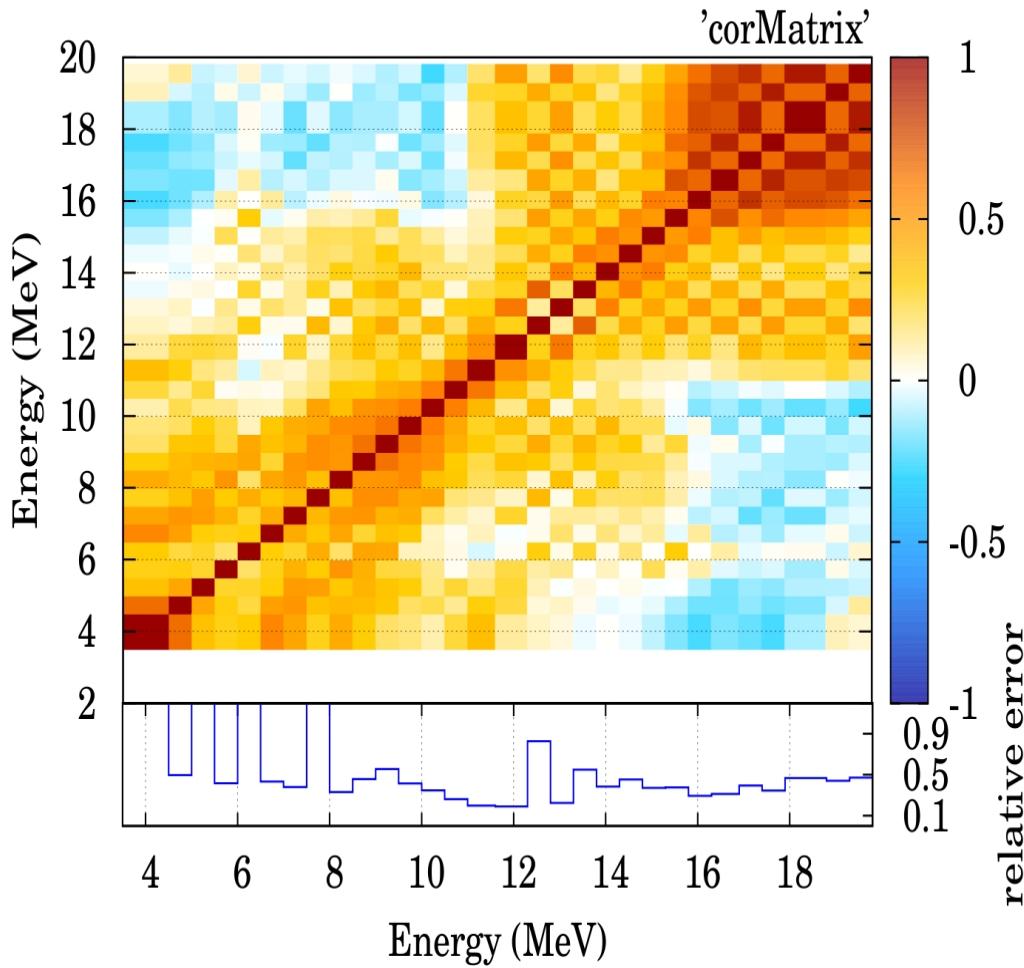
comparison of n,p



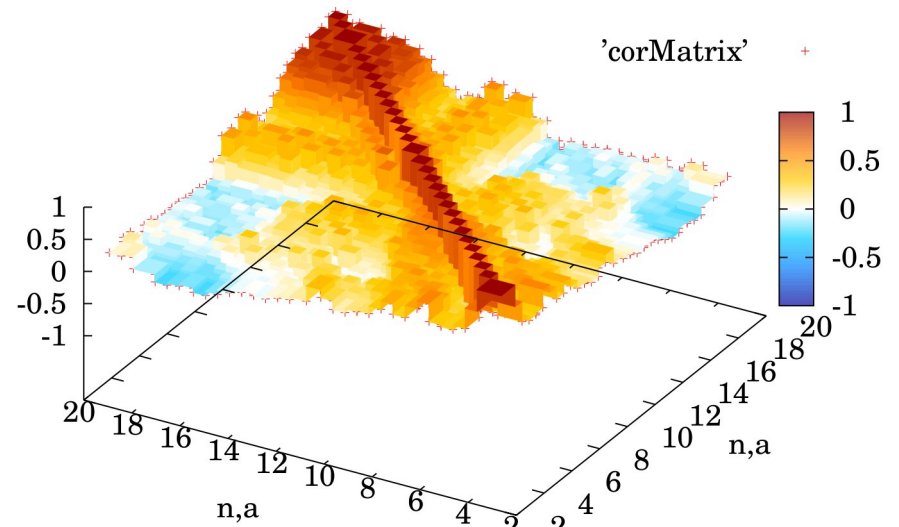
comparison of n, α



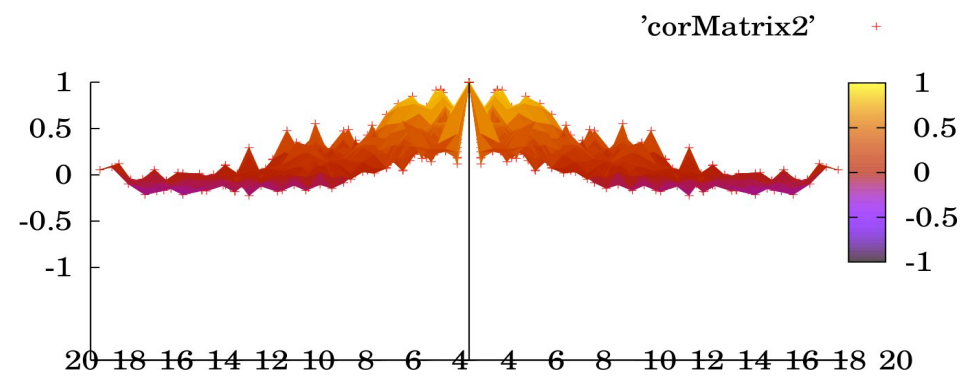
correlation (n,a) with (n,a)

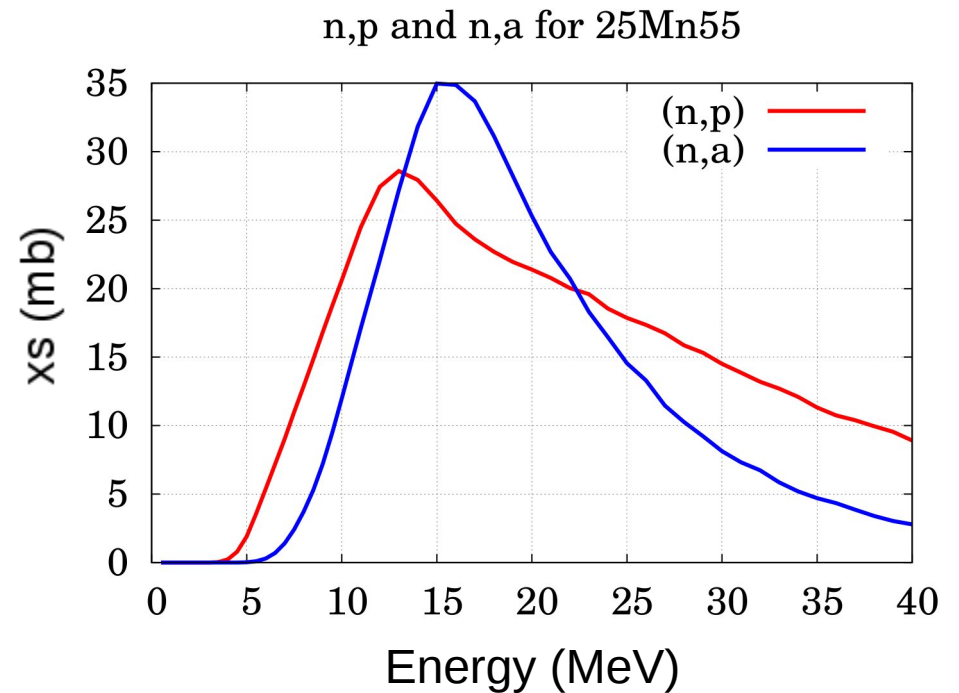
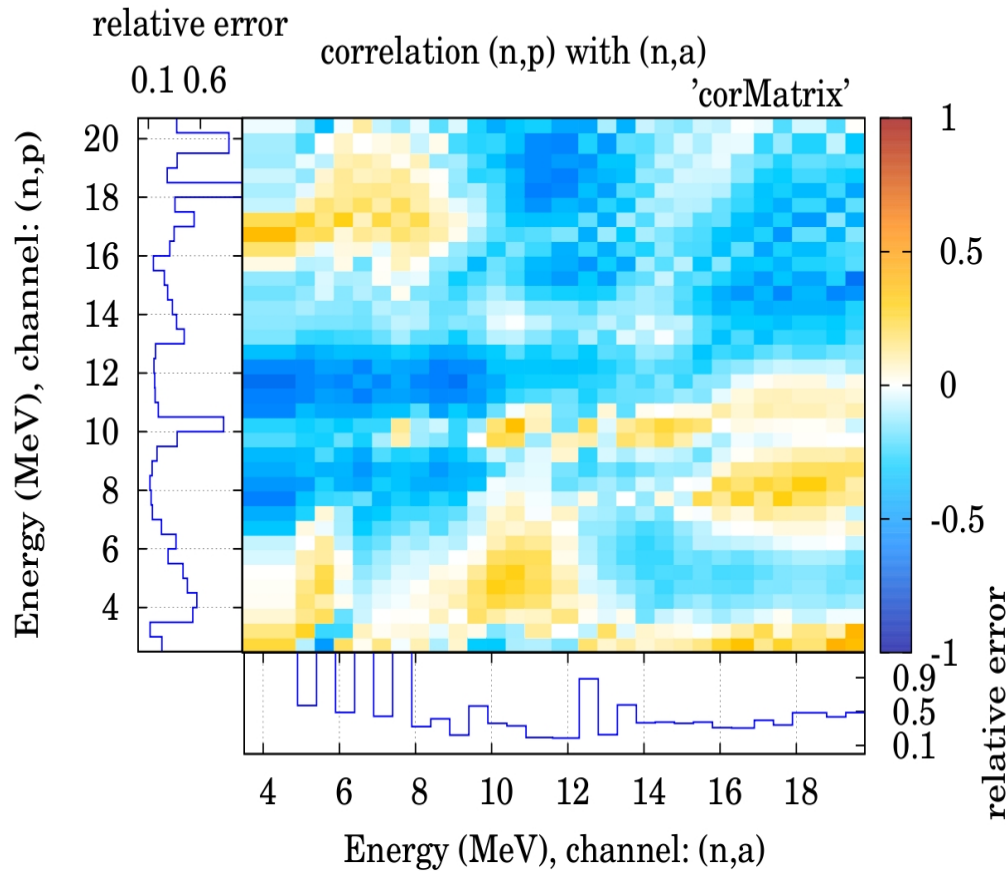


correlation matrix n,a with n,a



correlation matrix n,a with n,a





-) calculation of parameter uncertainties are in progress
-) Full prior is starting point for Bayesian update procedure together with exp. information of Mn55
-) errors will go down, if good experimental data available
-) evaluation of ^{55}Mn which will be performed in our workgroup under the supervision of Prof. H. Leeb in Vienna

Thank you for your attention!